



Side-Channel Expectation-Maximization Attacks

CHES 2022 Leuven

Julien Béguinot^{1,2} Wei Cheng^{1,2} Sylvain Guilley^{1,2} Olivier Rioul¹

¹ LTCI, Télécom Paris, Institut Polytechnique de Paris

² Secure-IC SAS





A Talk About Side-Channel Analysis

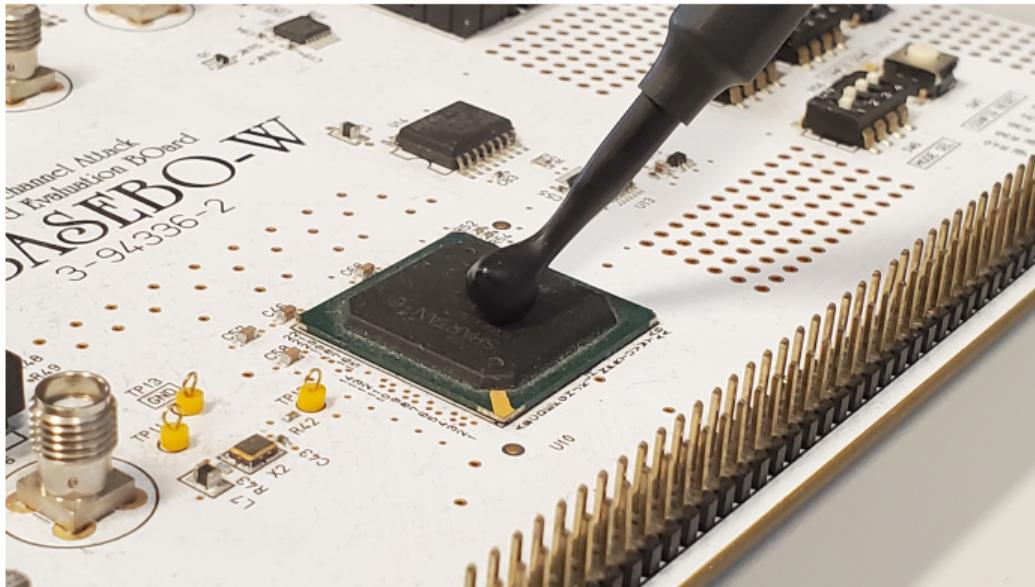


Figure: An Electromagnetic Capture for SCA



State of the Art

■ Supervised

■ Profiling Phase

■ Parametric Model

Properties	Supervised	Profiling Phase	Parametric Model	Flexible Model
Template [Chari et al.]	✓	✓	✓	✓
P-EM [Lemke et al.]	—	✓	✓	✓
Proposed U-EM [This talk]	—	—	✓	✓
2O-CPA [Prouff et al.]	—	—	✓	—
MIA/KSA [Whitnall et al.]	—	—	—	—

We bridge the following gap. How to leverage parametric attack in presence of masking without prior profiling ? In particular the proposed U-EM outperforms 2O-CPA .

Notations and Setup

- An n -bit secret key k encrypts Q plaintext bytes $\mathbf{t} = (t_1, t_2, \dots, t_Q)$
- The random masks $\mathbf{M} = (M_1, M_2, \dots, M_Q)$ are i.i.d. uniform in \mathbb{F}_2^n
- The sensitive variable $\mathbf{X} = (X_1, X_2, \dots, X_Q)$ is such that

$$X_q = S(k \oplus t_q) \oplus M_q \quad (q = 1, 2, \dots, Q) \quad (1)$$

where S is the substitution box

- The noise is assumed Additive White Gaussian (AWGN)
- A bivariate leakage model $\mathbf{Y} = (Y_1, \dots, Y_Q)$ where for each $q \in \{1, 2, \dots, Q\}$ one has $Y_q = (Y_q^{(1)}, Y_q^{(2)})$ with noise variance $\sigma^2 = (\sigma_1^2, \sigma_2^2)$

Leakage Models

■ Hamming Weight Leakage Model.

$$\begin{cases} Y_q^{(1)} = a^{*,(1)} w_H(X_q) + b^{*,(1)} + N_q^{(1)} \\ Y_q^{(2)} = a^{*,(2)} w_H(M_q) + b^{*,(2)} + N_q^{(2)} \end{cases} \quad (q = 1, 2, \dots, Q) \quad (2)$$

where $a^{*,(i)} \in \mathbb{R}$ and $b^{*,(i)} \in \mathbb{R}$ are unknown parameters.

■ Linear Leakage Model.

$$\begin{cases} Y_q^{(1)} = \langle a^{*(1)}, X_q \rangle + b^{*(1)} + N_q^{(1)} \\ Y_q^{(2)} = \langle a^{*(2)}, M_q \rangle + b^{*(2)} + N_q^{(2)} \end{cases} \quad (q = 1, 2, \dots, Q) \quad (3)$$

where $\langle \cdot, \cdot \rangle$ denotes a bitwise scalar product over the reals and where vector $a^{*(i)} \in \mathbb{R}^n$ and $b^{*(i)} \in \mathbb{R}$ are unknown parameters.

■ Quadratic Leakage Model.

More general leakage model, with interaction between bits:

$$\begin{cases} Y_q^{(1)} = X_q^\top a^{*(1)} X_q + b^{*(1)} + N_q^{(1)} \\ Y_q^{(2)} = M_q^\top a^{*(2)} M_q + b^{*(2)} + N_q^{(2)} \end{cases} \quad (q = 1, 2, \dots, Q), \quad (4)$$

Generic Notation for the Sensitive Variable

To simplify the derivations we introduce the notations.

■ Hamming Weight Model

$$\begin{cases} x^{(1)}(a, b, k, t, m) = a \cdot w_H(m) + b \\ x^{(2)}(a, b, k, t, m) = a \cdot w_H(S(k \oplus t) \oplus m) + b \end{cases} \quad (5)$$

■ Linear Model

$$\begin{cases} x^{(1)}(a, b, k, t, m) = \langle a, m \rangle + b \\ x^{(2)}(a, b, k, t, m) = \langle a, S(k \oplus t) \oplus m \rangle + b \end{cases} \quad (6)$$

■ Quadratic Model

$$\begin{cases} x^{(1)}(a, b, k, t, m) = m^\top a m + b \\ x^{(2)}(a, b, k, t, m) = (S(k \oplus t) \oplus m)^\top a (S(k \oplus t) \oplus m) + b \end{cases} \quad (7)$$



Theoretical Optimal Distinguisher [Heuser et al.]

The Maximum Likelihood (ML)-based distinguisher is

$$\hat{k}(y) = \arg \max_k \sum_{q=1}^Q \log \left[\sum_{m_q \in \mathbb{F}_2^n} \exp \left(-\frac{1}{2} \|y_q - x(a^*, b^*, k, t_q, m_q)\|^2 \right) \right]. \quad (8)$$

The template based distinguisher [Chari et al.] uses the ML-based distinguisher with estimation \hat{a}, \hat{b} of the parameters a^*, b^* .

2O-CPA [Prouff et al.]

For a given key hypothesis k , we write $x(k) = (X(k)_1, \dots, X(k)_Q)$ where

$$x(k)_q = \frac{1}{2^n} \sum_{m \in \mathbb{F}_2^n} w_H(m) w_H(S(k \oplus t_q) \oplus m) \quad (q = 1, 2, \dots, Q). \quad (9)$$

The distinguisher is then

$$\hat{k}_{\text{2O-CPA}}(y) = \arg \max_k |\rho(x(k), \widetilde{y^{(1)}} \widetilde{y^{(2)}})| = \arg \max_k \left| \frac{\text{Cov}(x(k), \widetilde{y^{(1)}} \widetilde{y^{(2)}})}{\sigma_{x(k)} \sigma_{y^{(1)} y^{(2)}}} \right| \quad (10)$$

where ρ is the empirical Pearson correlation coefficient and \widetilde{y} denotes the centered version of the vector y .

Expectation Maximization for SCA

$$a_{p+1}, b_{p+1} \leftarrow \underbrace{\arg \max_{a,b}}_{\text{M-Step}} \underbrace{\mathbb{E}_{M \sim \mathcal{U}(\mathbb{F}_2^n)^Q} [\log(\mathbb{P}(Y = y, M | k, a, b))]}_{\text{E-Step}} \quad (11)$$



Bayes Posterior

For fixed k and q , let

$$\beta_q^{(p)}(m) = \mathbb{P}(M_q = m) \exp\left(-\frac{1}{2} \|y_q - x(a_p, b_p, k, t_q, m)\|^2\right).$$

Then the Bayes posterior of the mask m being used for the q -th traces given a_p, b_p is

$$\alpha_q^{(p)}(m)(= \mathbb{P}(M_q = m | y, a_p, b_p)) = \frac{\beta_q^{(p)}(m)}{\sum_{m'} \beta_q(m')^{(p)}}.$$

Explicit E-Step

The E-Step at the p -th iteration is derived as follows.

$$\begin{aligned}\mathbb{E}[\log(\mathbb{P}(Y = y, M|a, b))] &= \mathbb{E} \left[\sum_q \log(\mathbb{P}(Y_q = y_q, M_q|a, b)) \right] \\ &= \sum_q \sum_{m_q} \mathbb{P}(M_q = m_q | Y_q = y_q, a_p, b_p) \log(\mathbb{P}(Y_q = y_q, M_q = m_q | a, b)) \\ &= \sum_{q,m} \alpha_q^{(p)}(m) \log(\mathbb{P}(Y_q = y_q | M_q = m, a, b)) + cst\end{aligned}\quad (12)$$

where the constant cst is independent of a and b . Thus the E-Step of the EM reduces to

$$(a_{p+1}, b_{p+1}) \leftarrow \arg \min_{(a,b)} \sum_q \sum_{m_q} \alpha_q^{(p)}(m_q) \|y_q - x(a, b, k, t_q, m_q)\|^2. \quad (13)$$

M-Step for Hamming Weight Model

Then the empirical covariance and variance are:

$$\widehat{\text{Cov}_{xy}^{(i)}} = \frac{1}{Q} \sum_{q,m} \alpha_q(m) y_q^{(i)} x_{k,t_q,m}^{(i)}$$
$$\widehat{\text{Var}_x^{(i)}} = \frac{1}{Q} \sum_{q,m} \alpha_q(m) y_q^{(i)2}$$

The M-Step is given by the following update rule ($i = 1, 2$):

$$a^{(i)} = \frac{\widehat{\text{Cov}_{xy}^{(i)}}}{\widehat{\text{Var}_x^{(i)}}} \quad \text{and} \quad b^{(i)} = -a^{(i)} \bar{x}^{(i)} \quad (14)$$

M-Step for Linear Model

The empirical autocorrelation matrix and intercorrelation vector are

$$\widehat{R}_{xx}^{(i)} = \sum_{q,m} \alpha_q(m) (x_{k,t_q,m}^{(i)} - \bar{x}^{(i)}) (x_{k,t_q,m}^{(i)} - \bar{x}^{(i)})^\perp \in \mathbb{R}^{n \times n}.$$

$$\widehat{R}_{xy}^{(i)} = \sum_{q,m} \alpha_q(m) (x_{k,t_q,m}^{(i)} - \bar{x}^{(i)}) y_q^{(i)\perp} \in \mathbb{R}^n$$

The M-Step is given by the following update rule ($i = 1, 2$):

$$a^{(i)} = (\widehat{R}_{xx}^{(i)})^{-1} \widehat{R}_{xy}^{(i)} \quad \text{and} \quad b^{(i)} = -\langle a^{(i)}, \bar{x}^{(i)} \rangle \quad (15)$$

For the quadratic model the M-Step is performed with a gradient based optimizer.

Pseudo-Code of U-EM

Algorithm 2: Pseudo-code: U-EM-LIN.

Data: The traces $\mathbf{y} = (y_1, \dots, y_Q)$ and the noise standard deviation
 $\sigma = (\sigma^{(1)}, \sigma^{(2)})$

Input: Convergence threshold ϵ

Output: Estimated key \hat{k}

```
1  $\bar{\mathbf{y}} \leftarrow \frac{1}{Q} \sum_{q=1}^Q y_q ;$  // Precompute the mean of the traces
2  $y^{(i)} \leftarrow \frac{y^{(i)} - \bar{\mathbf{y}}^{(i)}}{\sigma^{(i)}}$  for  $i = 1, 2 ;$  // Center and normalise the traces by  $\sigma$ 
3 forall key hypothesis  $k \in \mathbb{F}_2^n$  do
    /* Initializations of the parameters  $a$  and  $b$  */  

    4  $a, b \leftarrow ((1, \dots, 1), (1, \dots, 1)), (0, 0) ;$  // Arbitrary, but could be chosen
    5 while True do
```

```

/* E-Step */
forall  $q, m$  compute do
   $x_{k,t_q,m} \leftarrow (S(k \oplus t_q) \oplus m, m)$ 

forall  $q$  do
   $c_q \leftarrow \max_m -\frac{1}{2} \|y_q - \langle a, x_{k,t_q,m} \rangle - b\|^2$ 

forall  $q, m$  do
   $\beta_q(m) \leftarrow p(m) \exp(c_q - \frac{1}{2} \|y_q - \langle a, x_{k,t_q,m} \rangle - b\|^2)$ 

forall  $q$  compute do
  forall  $m$  compute do
     $\alpha_q(m) \leftarrow \frac{\beta_q(m)}{\sum \beta_q(m')}$ 
     $\tilde{x}_q \leftarrow \sum_m \alpha_q(m) x_{k,t_q,m}$ 

   $\bar{x} \leftarrow \frac{1}{Q} \sum_{q=1}^Q \tilde{x}_q$ 

```

18

```
/* M-Step
for i =1,2 do
```

19

$$\widehat{R}_{\mathbf{xx}}^{(i)} \leftarrow \sum_{q,m} \alpha_q(m) (x_{k,t_q,m}^{(i)} - \bar{\mathbf{x}}^{(i)}) (x_{k,t_q,m}^{(i)} - \bar{\mathbf{x}}^{(i)})^\perp \in \mathbb{R}^{n \times n}$$

20

$$\widehat{R}_{\mathbf{xy}}^{(i)} \leftarrow \sum_q \alpha_q(m) (\tilde{x}_q^{(i)} - \bar{\mathbf{x}}^{(i)}) y_q^{(i)\perp} \in \mathbb{R}^n$$

21

```
if ( $\|a - a'\|^2 + \|b - b'\|^2$ ) <  $\epsilon$  then
```

22

```
Break ; // Exit condition
```

23

```
a,b  $\leftarrow a',b'$ 
```

24

$$LogLikelihood(k) \leftarrow \sum_{q=1}^Q \log(\sum \beta_q) + c_q$$

Result: $\hat{k} = \arg \max_k LogLikelihood(k)$



Profiled EM [Lemke-Rust et al.]

- Use EM to profile a template for each key Hypothesis.
- Require to have the necessary traces for profiling for each key hypothesis.
- Perform Maximum Likelihood with the derived templates.

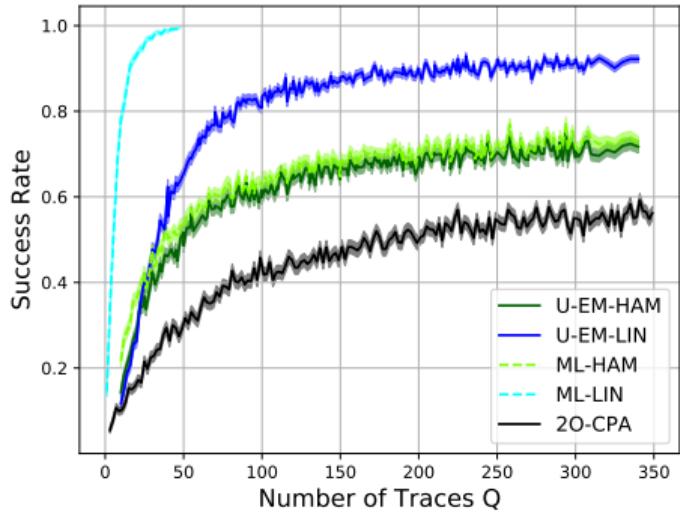


Numerical Evaluation Framework and Epistemic Noise

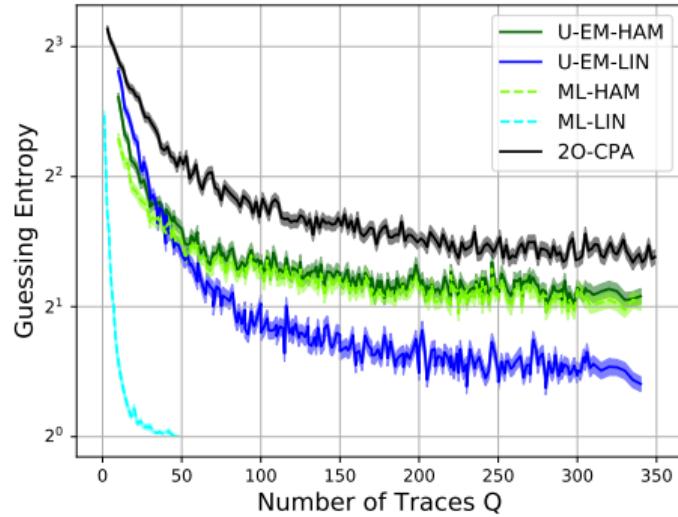
- 4 bits PRESENT substitution box.
- Noise with standard deviation σ
- *Epistemic Noise* with standard deviation σ_a
- 10^3 independant repetitions. All attacks are performed on the same traces.
- Real Data: DPA Contest V4.2 with 8 bits Sub-Byte look up table

The epistemic noise is AWGN on the leakage coefficients and represents the physical peculiarities of a leaking device.

Numerical Evaluations



(a) Success rate.



(b) Guessing entropy.

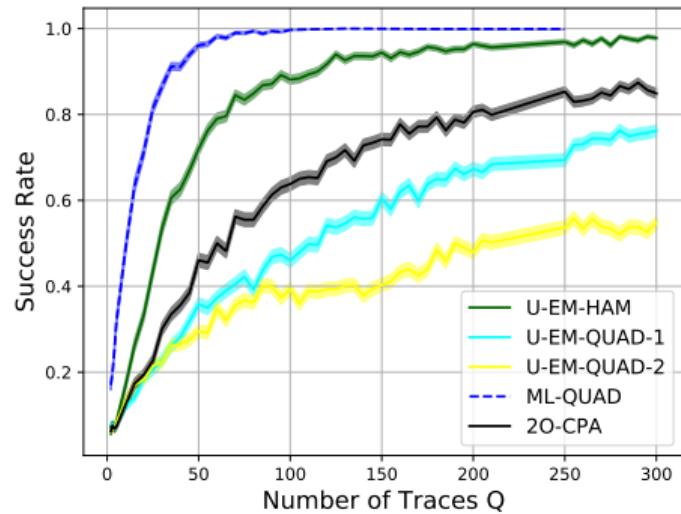
Figure: Attack metrics evaluated with $\sigma = 0.3$ and $\sigma_a = 0.8$.



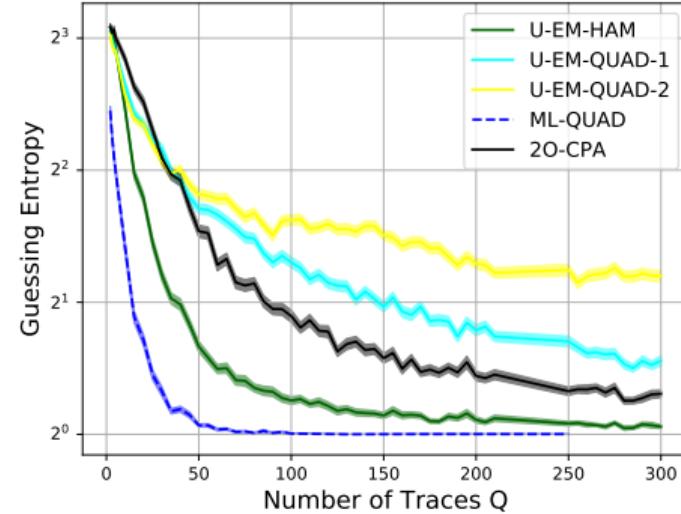
Tikhonov regularization in M-Step

The model complexity can be a burden for the U-EM. Hence it can be interesting to consider the well-known Tikhonov regularization (a.k.a ridge regression) as already suggested by [Wang et al.]. In this case we add this regularization term in the M -step.

Numerical Evaluations with Quadratic Model



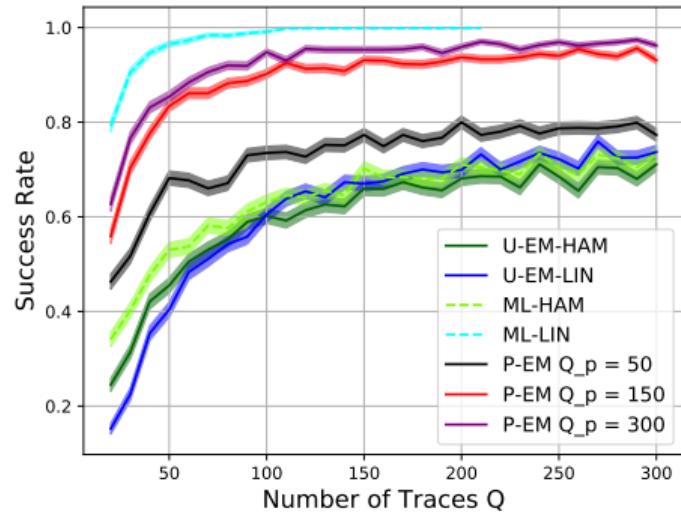
(a) Success rate.



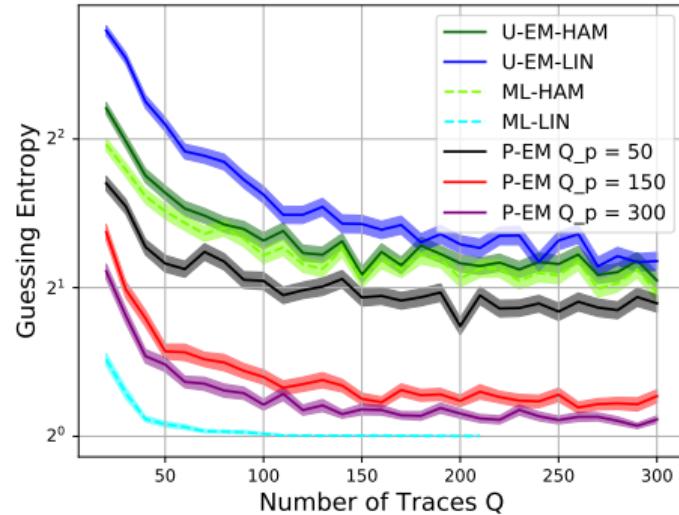
(b) Guessing entropy.

Figure: Attack metrics evaluated with $\sigma = 1$ and $\sigma_a = 0.8$.

Comparison with P-EM



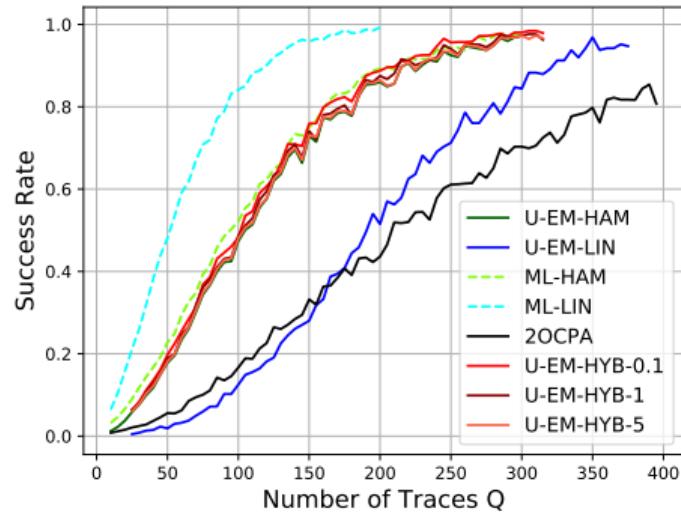
(a) Success rate on the DPA Contest.



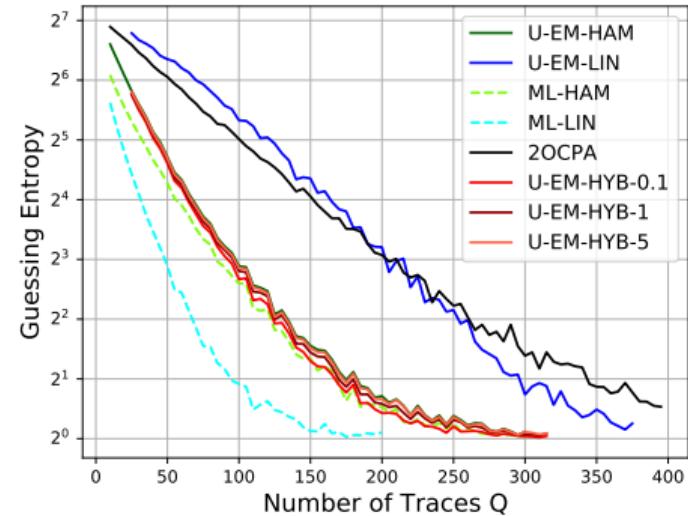
(b) Guessing entropy on the DPA Contest.

Figure: Attack metrics evaluated with $\sigma = 0.4$ and $\sigma_a = 0.4$.

Results on the Real Data from DPA Contest



(a) Success rate on the DPA Contest.



(b) Guessing entropy on the DPA Contest.

Figure: Attack metrics on the traces of the DPA Contest.



Conclusion and Perspectives

- U-EM is interesting when profiling is not feasible
- U-EM is computationally efficient
- U-EM can be easily adapted to higher order attacks and is flexible
- When profiling is possible template or P-EM are better than our U-EM as it suffers from "overfitting"
- U-EM fails to generalize to the quadratic model. Can unprofiled attack handle more complex model ? Choose best leakage basis ?



References

- Weijia Wang, Yu Yu, François-Xavier Standaert, Junrong Liu, Zheng Guo, and Dawu Gu, *Ridge-Based DPA: Improvement of Differential Power Analysis For Nanoscale Chips*, IEEE Trans. 2018
- Suresh Chari, Josyula R. Rao, and Pankaj Rohatgi, *Template attacks*, CHES 2002
- Benedikt Gierlichs, Lejla Batina, Pim Tuyls, and Bart Preneel, *Mutual information analysis*, CHES
- Kerstin Lemke-Rust and Christof Paar, *Gaussian mixture models for higher-order side channel analysis*, CHES 2007
- Houssem Magharebi, Claude Carlet, Sylvain Guilley, and Jean-Luc Danger, *Optimal first-order masking with linear and non-linear bijections*, AFRICACRYPT 2012
- Emmanuel Prouff, Matthieu Rivain, and Régis Bevan, *Statistical Analysis of Second Order Differential Power Analysis*, IEEE Trans. Computers 58 (2009)
- Carolyn Whitnall, Elisabeth Oswald, and Luke Mather, *An Exploration of the Kolmogorov-Smirnov Test as a Competitor to Mutual Information Analysis*, CARDIS



Thank You For Your Attention

CHES 2022 Leuven

Julien Béguinot^{1,2} Wei Cheng^{1,2} Sylvain Guilley^{1,2} Olivier Rioul¹

¹ LTCI, Télécom Paris, Institut Polytechnique de Paris

² Secure-IC SAS

